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Information borrowing in Bayesian clinical trials: choice of tuning parameters for the robust mixture prior

Vivienn Weru

Joint work with: Annette Kopp-Schneider, Manuel Wiesenfarth, Sebastian Weber & Silvia Calderazzo

PSI 2026: Belfast

Borrowing external data

- Bayesian clinical trial designs increasingly used:
 - save resources
 - timely delivery of new treatments to patients
- Borrowing external data
- Bayesian methods provide a natural way: use of informative priors
- Challenge: heterogeneity between external and current studies (aka **prior-data conflict**)



Dynamic borrowing

- Prior-data conflict remedy: Dynamic borrowing
- Borrow most when the current and external data are observed to be similar and least otherwise
- Examples of such methods:
 - Robust mixture prior (see Schmidli et al., 2014)
 - Power prior (see Ibrahim and Chen, 2000)
 - Commensurate prior (see Hobbs et al., 2012)
 - Compromise decision (see Calderazzo et al., 2024)
 -

All these methods depend on choices of borrowing parameters that influence their behaviour.

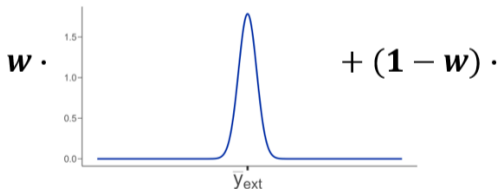
Robust mixture prior

For a parameter of interest θ , the mixture prior is given as follows (see Schmidli et al., 2014) :

$$\pi_{\text{mixture}}(\theta) = w\pi_{\text{ext}}(\theta) + (1 - w)\pi_{\text{robust}}(\theta)$$

Informative component (External data)

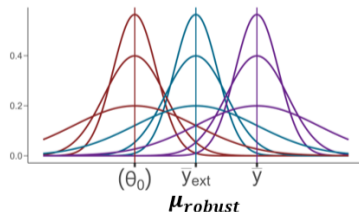
$$\pi_{\text{ext}}(\bar{y}_{\text{ext}}, \sigma_{\text{ext}}^2)$$



+ (1 - w) ·

Robust component

$$\pi_{\text{robust}}(\mu_{\text{robust}}, \sigma_{\text{robust}}^2)$$



Posterior also a mixture: $\pi_{\text{mixture}}(\theta|y) = \tilde{w}\pi_{\text{ext}}(\theta|y) + (1 - \tilde{w})\pi_{\text{robust}}(\theta|y)$

Scoping review

- Scoping review about use of robust Bayesian methods in Phase II trials (submitted)
- Robust mixture prior among the most used methods
- In some cases,
 - rationale for the choice of prior weight missing/not clear
 - exact prior specification of robust component not reported
- Poor reporting of borrowing parameters in general

Preprint



One-arm trial

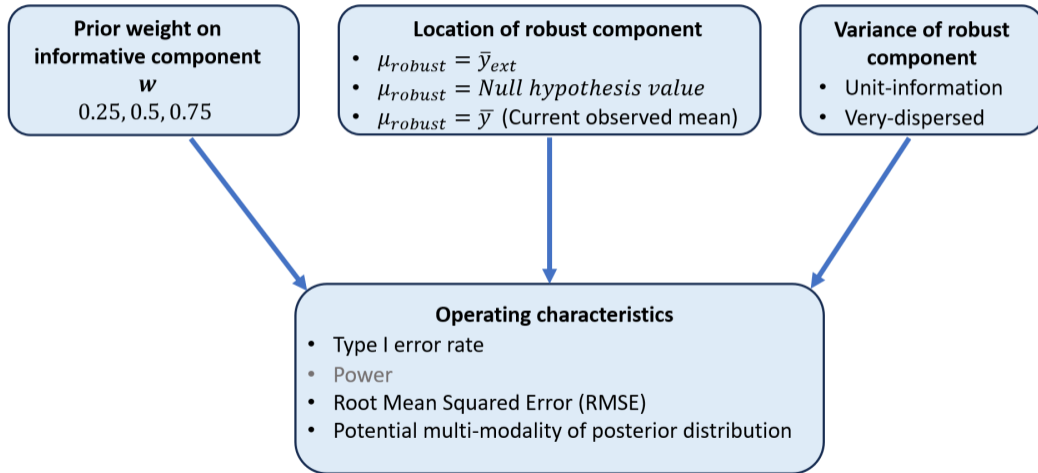
External
patients



Current
patients



Impact of robust component's parameter choices (Weru et al., 2026)



Set-up ($H_0 : \theta \leq 0$ vs $H_1 : \theta > 0$)

- Normal endpoint with known variance, $\sigma^2 = 1$:

$$\bar{y} \sim N(\theta, 1/n_{\text{current}}), \quad n_{\text{current}} = 20$$

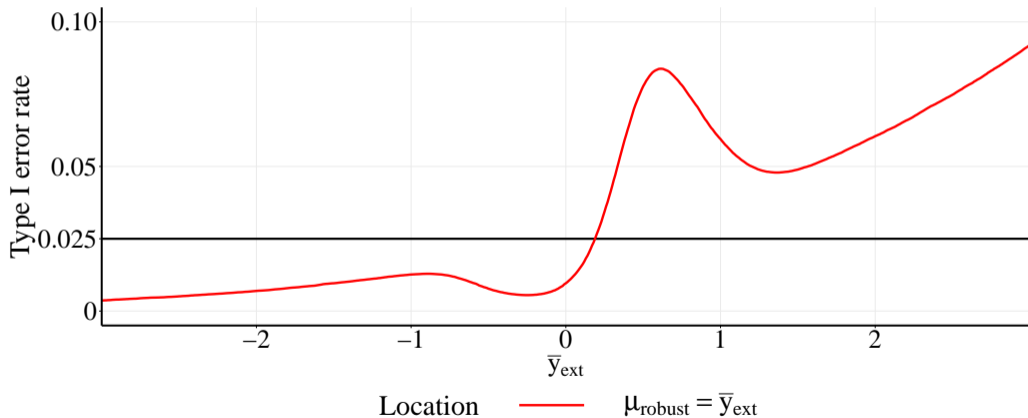
Mixture prior:

$$\begin{aligned}\pi_{\text{ext}}(\theta) &\sim N(\mu_{\text{ext}}, 1/n_{\text{ext}}), \\ \pi_{\text{robust}}(\theta) &\sim N(\mu_{\text{robust}}, 1/n_{\text{robust}}) \\ n_{\text{ext}} &= 15 \quad \text{and} \quad \text{unit-information:} \quad n_{\text{robust}} = 1\end{aligned}$$

- Since we evaluate Type I error, to generate the data $\theta = \theta_0 = 0$.

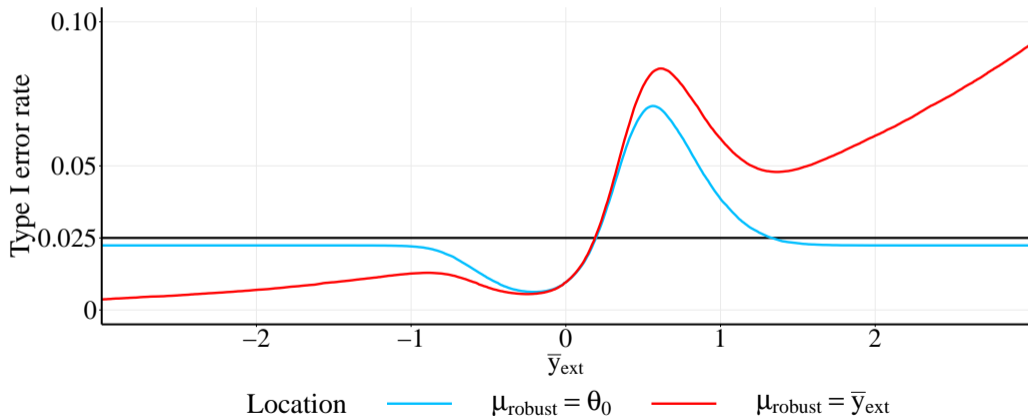
Testing θ ($H_0 : \theta \leq 0$ vs $H_1 : \theta > 0$): Type I error rate

$\pi_{\text{robust}}(\theta) \sim N(\mu_{\text{robust}}, 1/n_{\text{robust}})$, unit-information: $n_{\text{robust}} = 1$, $w = 0.5$



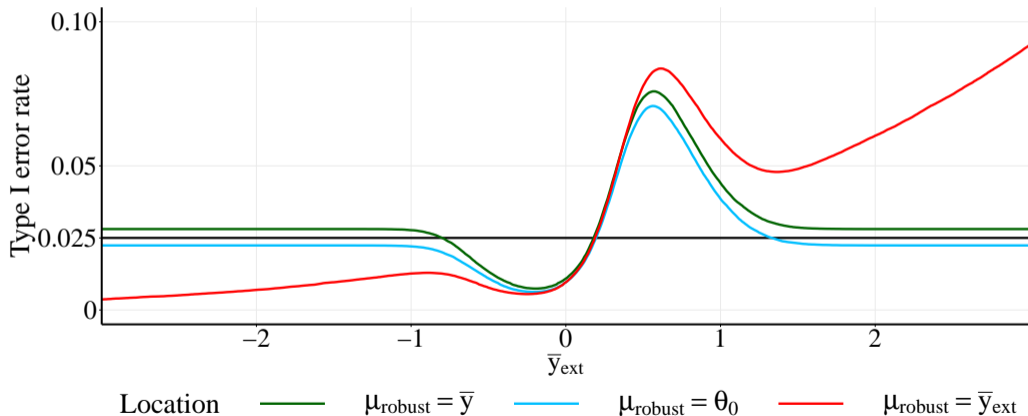
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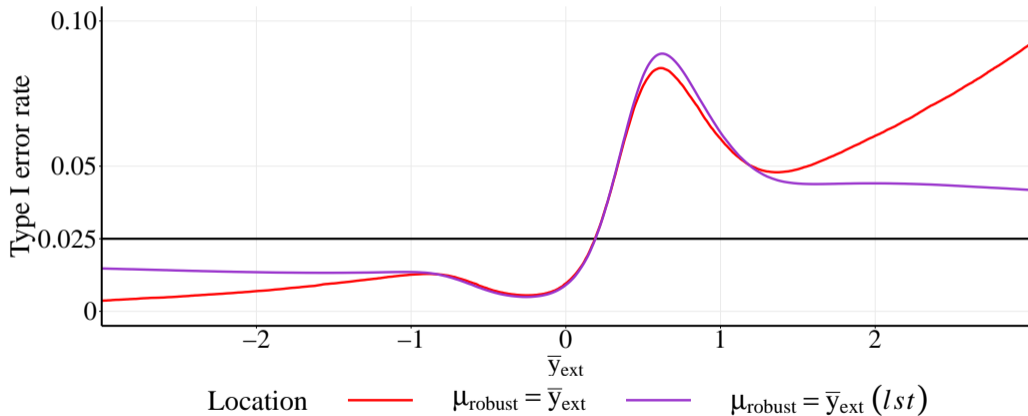
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t-distribution for robust component

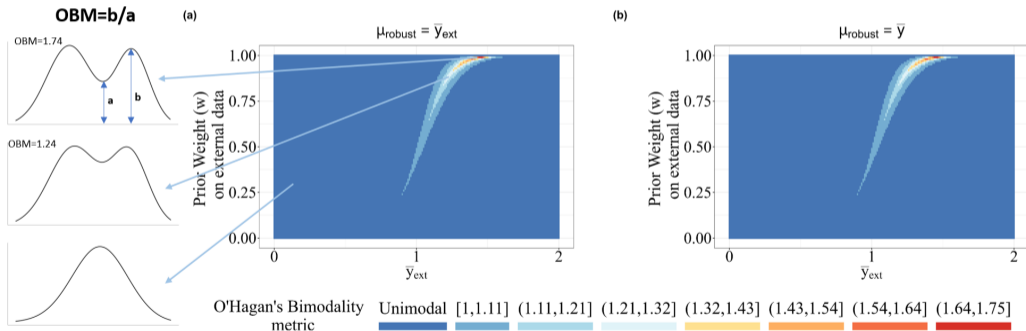
$\pi_{\text{robust}}(\theta) \sim \text{lst}(\mu, \sigma, \nu), \sigma = 1, \nu = 3, w = 0.5$ (*lst* short for location scale t distribution)



Requires MCMC.

Bimodality of posterior distribution

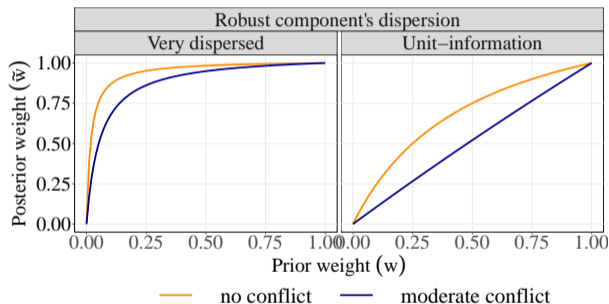
- Mixture prior \rightarrow Posterior also mixture
- O'Hagan's Bimodality Metric (OBM): minimum ratio of each mode's density to antimode's density (see O'Hagan and Forster, 2004)



Here, $\bar{y} = \theta$

Robust component's dispersion

- Robust component should not be too vague → high posterior weight given to informative component (even when there's conflict) (see e.g. Röver et al., 2019, Mutsvari et al., 2016)
- Unit-information prior has been proposed (Kass and Wasserman, 1995)



Also in the paper:

- Importance of robust component's location decreases with increasing dispersion.
- Bimodality more likely with less precise robust component.

Borrowing to control arm (Hybrid-control trials)

External control



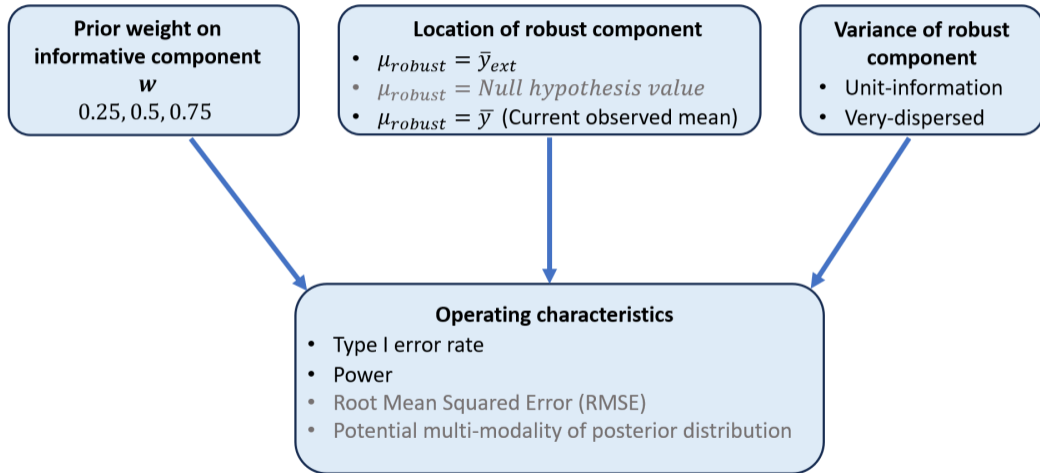
Current control



Current treated



Impact of robust component's parameter choices (Weru et al., 2026)



Set-up ($H_0 : \theta_t \leq \theta_c$ versus $H_1 : \theta_t > \theta_c$)

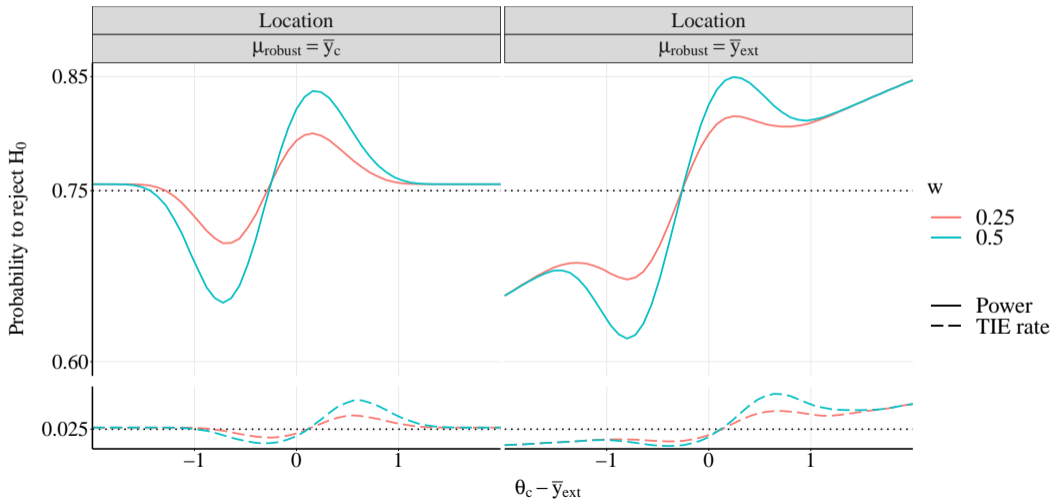
- External data available for the control arm
- Normal endpoint with known variance, $\sigma^2 = 1$:

$$y_t | \theta_t, \sigma_t \sim N(\theta_t, \sigma_t^2), \quad y_c | \theta_c, \sigma_c \sim N(\theta_c, \sigma_c^2), \quad \sigma_t^2 = \sigma_c^2 = \sigma^2$$
$$n_t = n_c = 20$$

Mixture prior:

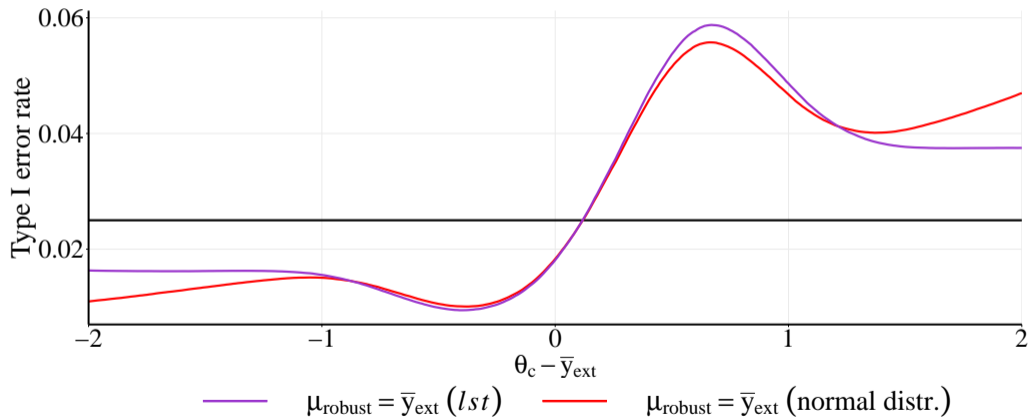
$$\pi_{\text{ext}}(\theta) \sim N(\mu_{\text{ext}}, 1/n_{\text{ext}}),$$
$$\pi_{\text{robust}}(\theta) \sim N(\mu_{\text{robust}}, 1/n_{\text{robust}})$$
$$n_{\text{ext}} = 15 \quad \text{and} \quad \text{unit-information:} \quad n_{\text{robust}} = 1$$

TIE rate and power



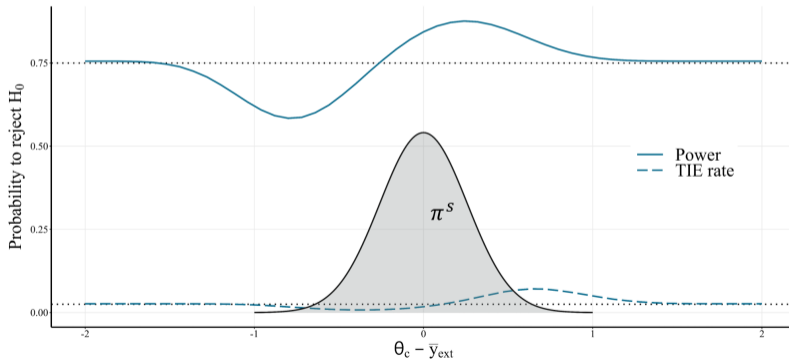
t-distribution for robust component

$\pi_{\text{robust}}(\theta) \sim \text{lst}(\mu, \sigma, \nu)$, $\sigma = 1, \nu = 3, w = 0.5$ (*lst* short for location scale t distribution)



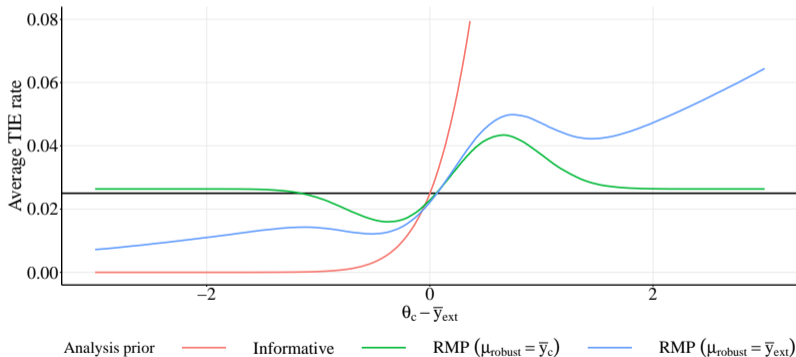
Average Type I error rate

- Bayesian metrics have been proposed (see e.g. Berry et al., 2010; Best et al., 2025; Calderazzo et al., 2020)
- Sampling/Design prior needed: used to generate a value of the parameter which is then used to generate the data



Average Type I error rate

- Informative component (external data) as design prior
- Different analysis priors



If design and analysis priors match, average type I error rate is controlled at 2.5% (Best et al., 2025).

Summary

- Parameters of robust component in mixture prior need careful thought.
- With small current sample size, impact of the unit-information prior on the posterior will be more pronounced.
- Location of robust component: most favourable behaviour depends on whether testing or estimation
 - Testing:
 - μ_{robust} at Null hypothesis value (one-arm trial)
 - μ_{robust} at current control observed mean (hybrid-control trial)
 - Estimation: $\mu_{\text{robust}} = \bar{y}$ (See Weru et al. (2026) for estimation framework)
- t-distribution more robust: choice of scale and degrees of freedom.
- Bimodality in the posterior distribution can arise and can complicate decision making.

Thank you!



Weru et al. (2026).
Statistics in Biopharmaceutical Research, 1-25

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Berry, S. M., Carlin, B. P., Lee, J. J., & Muller, P. (2010). Bayesian Adaptive Methods for Clinical Trials. CRC press. <https://doi.org/10.1201/ebk1439825488>



Best, N., Ajimi, M., Neuenschwander, B., Saint-Hilary, G., & Wandel, S. (2025). Beyond the classical type I error: Bayesian metrics for Bayesian designs using informative priors. Statistics in Biopharmaceutical Research, *17*(2), 183–196. <https://doi.org/10.1080/19466315.2024.2342817>



Calderazzo, S., Wiesenfarth, M., & Kopp-Schneider, A. (2020). A decision-theoretic approach to bayesian clinical trial design and evaluation of robustness to prior-data conflict. Biostatistics, *23*(1), 328–344. <https://doi.org/10.1093/biostatistics/kxaa027>



Calderazzo, S., Wiesenfarth, M., & Kopp-Schneider, A. (2024). Robust incorporation of historical information with known type I error rate inflation. Biometrical Journal, *66*(1), 2200322. <https://doi.org/10.1002/bimj.202200322>



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Ibrahim, J. G., & Chen, M.-H. (2000). Power prior distributions for regression models. Statistical Science, 46–60. <https://doi.org/10.1214/ss/1009212673>

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Mutsvari, T., Tytgat, D., & Walley, R. (2016). Addressing potential prior-data conflict when using informative priors in proof-of-concept studies. *Pharmaceutical Statistics*, *15*(1), 28–36. <https://doi.org/10.1002/pst.1802>



O'Hagan, A., & Forster, J. J. (2004). *Kendall's advanced theory of statistics, volume 2b: Bayesian inference (Vol. 2)*. Arnold.



Röver, C., Wandel, S., & Friede, T. (2019). Model averaging for robust extrapolation in evidence synthesis. *Statistics in Medicine*, *38*(4), 674–694. <https://doi.org/10.1002/sim.7991>



Schmidli, H., Gsteiger, S., Roychoudhury, S., O'Hagan, A., Spiegelhalter, D., & Neuenschwander, B. (2014). Robust meta-analytic-predictive priors in clinical trials with historical control information. *Biometrics*, *70*(4), 1023–1032. <https://doi.org/10.1111/biom.12242>

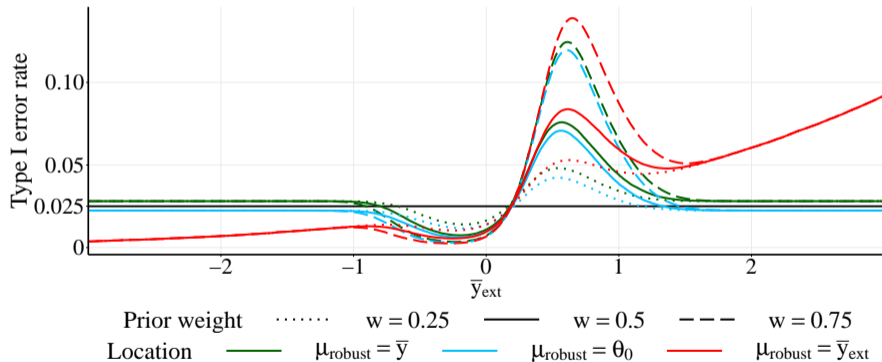


Weru, V., Kopp-Schneider, A., Wiesenfarth, M., Weber, S., & Calderazzo, S. (2026). Information borrowing in bayesian clinical trials: Choice of tuning parameters for the robust mixture prior. *Statistics in Biopharmaceutical Research*, 1–21. <https://doi.org/10.1080/19466315.2026.2646537>

Back-up slides

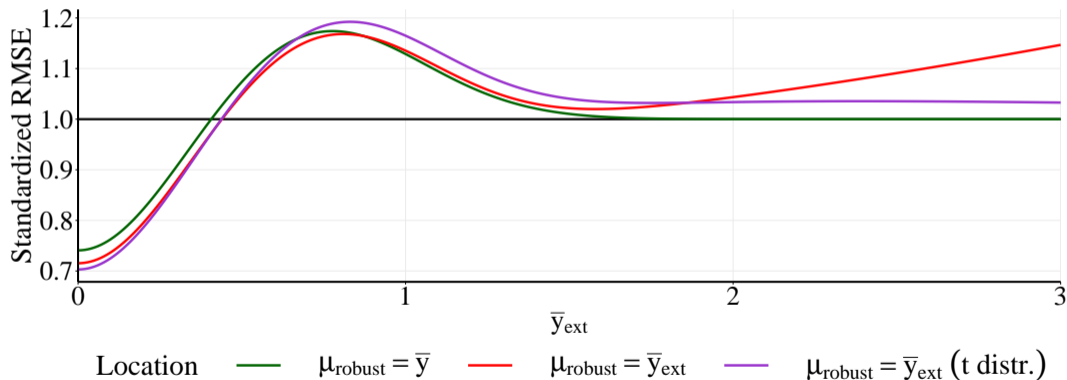
Testing θ ($H_0 : \theta \leq 0$ vs $H_1 : \theta > 0$): Type I error rate

$$\pi_{\text{robust}}(\theta) \sim N(\mu_{\text{robust}}, 1/n_{\text{robust}}), \text{ unit-information: } n_{\text{robust}} = 1$$



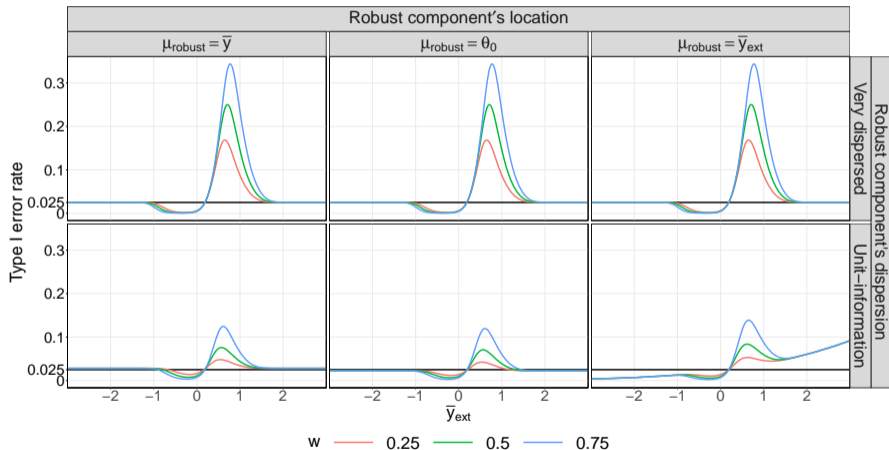
Estimating θ : Root Mean Squared error (RMSE)

$\pi_{\text{robust}}(\theta) \sim N(\mu_{\text{robust}}, 1/n_{\text{robust}})$, unit-information: $n_{\text{robust}} = 1$, $w = 0.5$



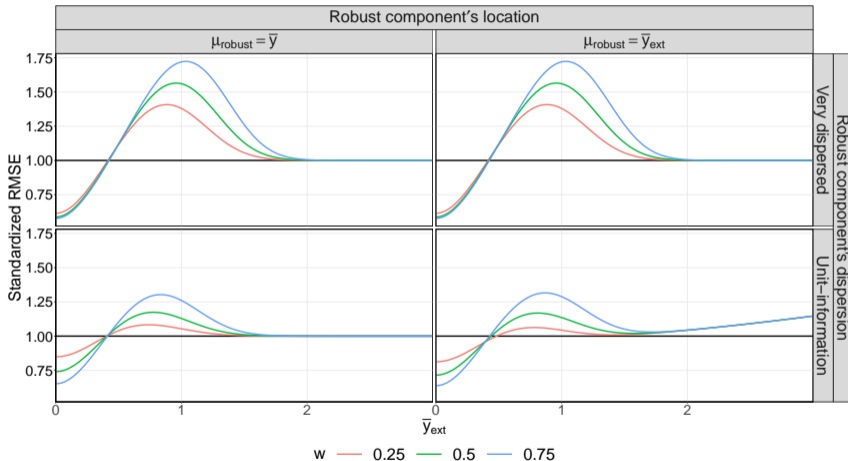
With $\mu_{\text{robust}} = \bar{y}$, RMSE better

Type I error rate: robust component's dispersion



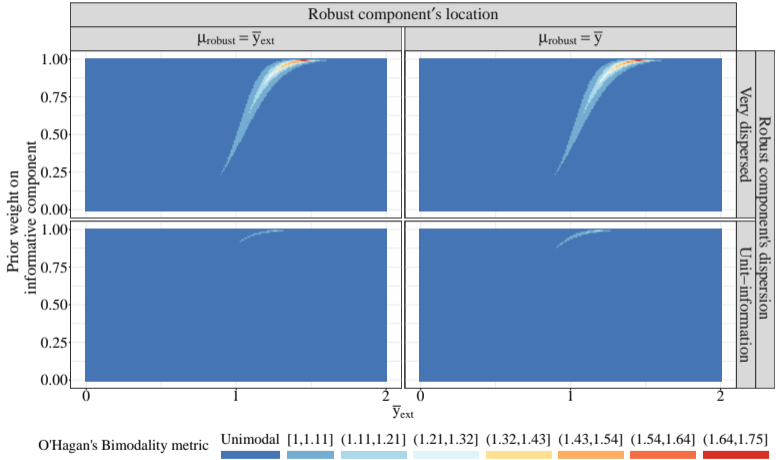
Importance of robust component's location decreases with increasing dispersion.

RMSE: robust component's dispersion



Location of robust component less important if component's precision is low

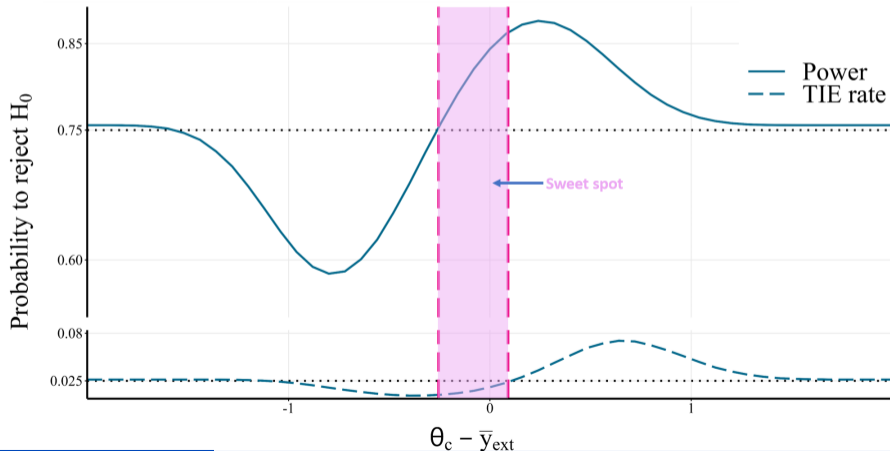
Bimodality: robust component's dispersion



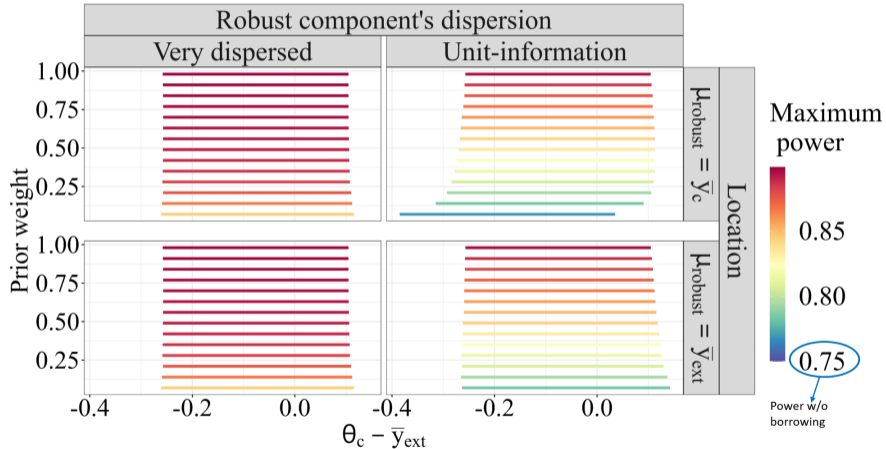
Bimodality more likely with less precise robust components and when μ_{ext} and μ_{robust} differ.

Sweet spot

In the hybrid-control setting, a "sweet spot" exists where gains in power and reductions in TIE rate are possible



Sweet spot



Robust component parameter choices impact length and position of sweet spot