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# Impact of interval-censored data on comparative time-to-event endpoints: a simulation study applied to patient-reported outcomes in oncology

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# Background

- > Patient-reported outcome (PRO) time-to-event endpoints feature as key endpoints in **SISAQOL-IMI recommendations**
- > In oncology clinical trials, a research question may be to evaluate whether patients in the treatment arm report **delayed PRO worsening**: time to deterioration (**TTD**)
- > PRO assessments are commonly assessed at scheduled visits, so true event times are only known to lie between visits: **interval-censored data**.
- > This interval censoring is not accounted for when using **Cox proportional hazards (PH)** or **Kaplan-Meier (KM) methods**, and its impact on treatment effect estimation is not routinely quantified

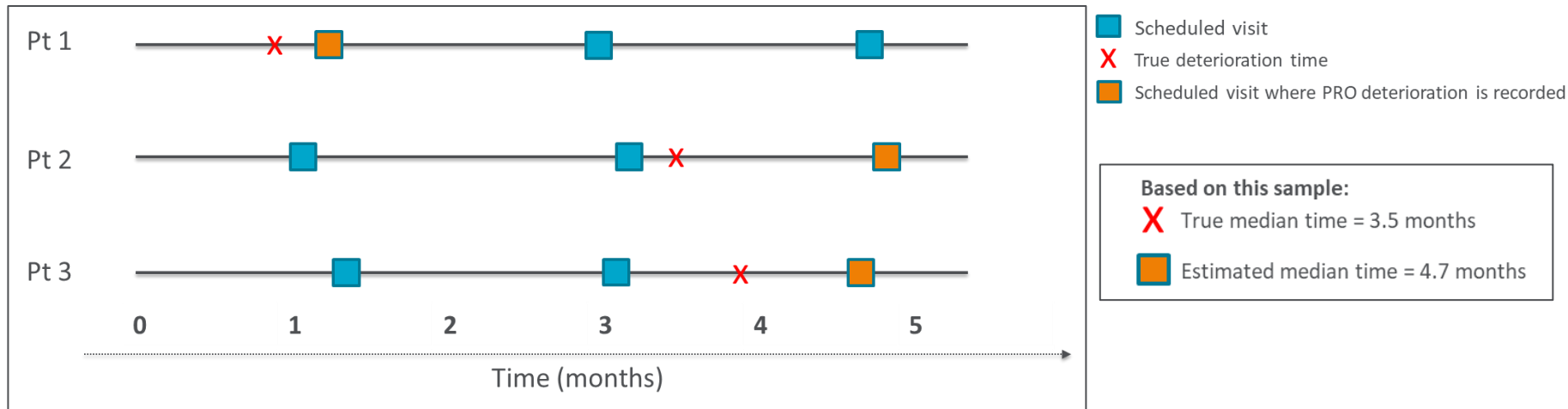
# TTD Literature Review

Initial on-treatment assessment schedule	Number of trials (% of N = 70)
Weekly	4 (5.7%)
Every 2 weeks	3 (4.3%)
Every 3 weeks	18 (25.7%)
Every 4 weeks	17 (24.3%)
Every 6 weeks	10 (14.3%)
Every 8 weeks	10 (14.3%)
Every 12 weeks	6 (8.6%)
Every 16 weeks	2 (2.9%)

TTD analysis methods	
Cox PH model	60 (85.7%)
Log-rank test	45 (64.3%)
Both Cox PH and log-rank	39 (55.7%)
Competing risk model	1 (1.4%)
Not reported	3 (4.3%)

# Why is interval censoring a problem?

- > **Fixed-visit imputation** – assigning the event time to the visit at which it is first detected



- > Other methods have been suggested:
  - **Mid-point imputation** – event time taken to be the mid-point of the interval
  - Other more sophisticated methods such as **Turnbull's estimator** (survival estimation) and **Finkelstein's Generalization of Cox Regression** (for effect estimation)

# Objectives

We simulated various realistic scenarios to investigate the following:

1. How does interval-censored data impact on **bias** for PRO TTD endpoints? Assuming **assessment frequency same** between arms.
2. What is the impact of **reduced/altered** assessment frequency following occurrence of intercurrent events (ICE)?

We focus on the **bias** of **Cox proportional hazards regression model** for estimating the hazard ratio (HR).

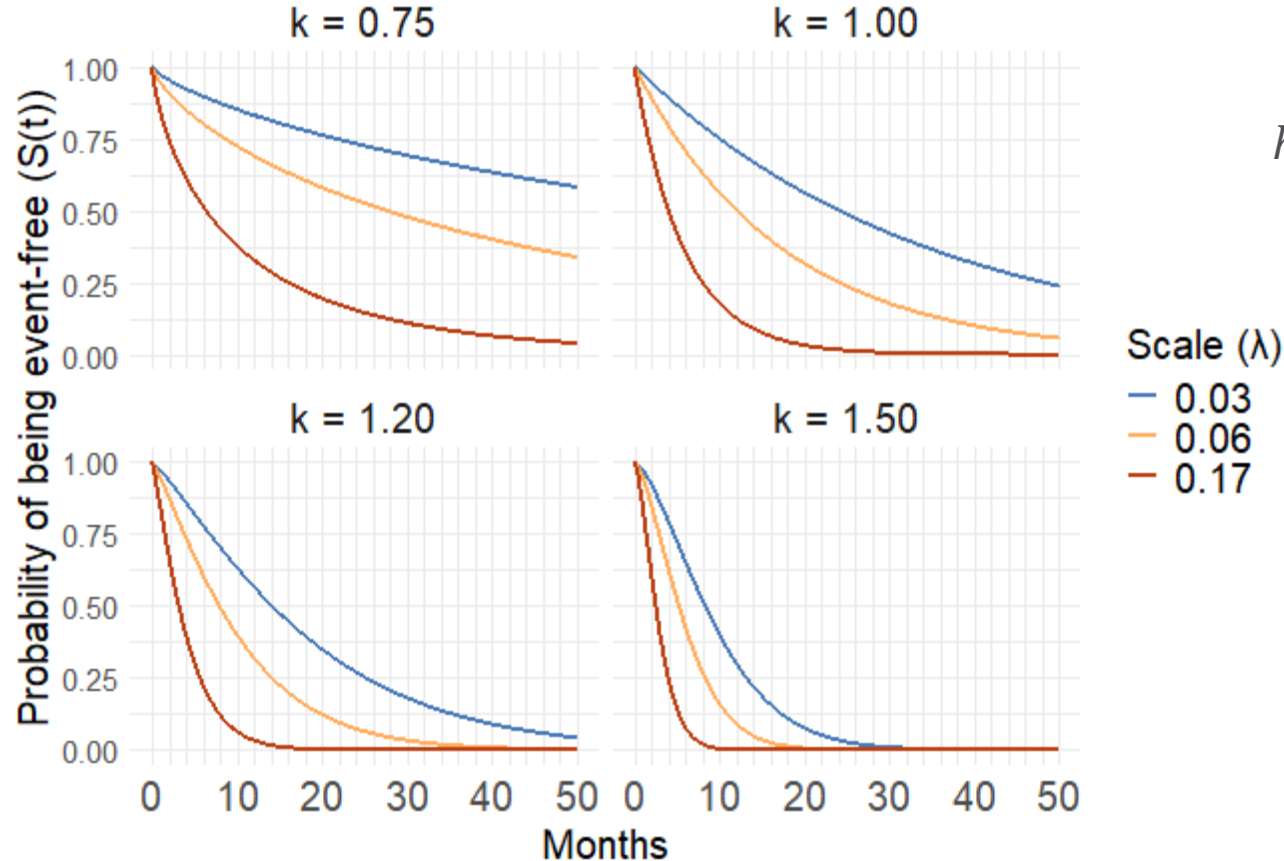
# Simulation methods: study design

- > 1:1 randomization ratio for two treatment groups,  $Trx = \{0,1\}$ 
  - Sample size ( $n$ ) = 400
- > Visits specified according to an assessment schedule with **fixed interval length  $\Delta$**  varied:  $\Delta = \{4, 8, 12\}$  weeks
  - Small **scheduling variability** for patient  $i$  and interval  $j$ :  $\Delta + \varepsilon_{ij}$ ,
  - Assume interval length **non-informative** and all assessments completed



# Simulation methods: true event times

Weibull Survival Functions for Different Shape (k) and Scale (λ) Parameter



> Assume **Cox PH** model

$$h(t) = h_0(t) \exp(\beta_{Trx} \times Trx)$$

> Simulate true event times based on **Weibull** distribution

> Multiple **baseline hazard shapes** assumed to mimic real PRO deterioration event time distributions

$$h_0(t) = \lambda k t^{k-1}$$

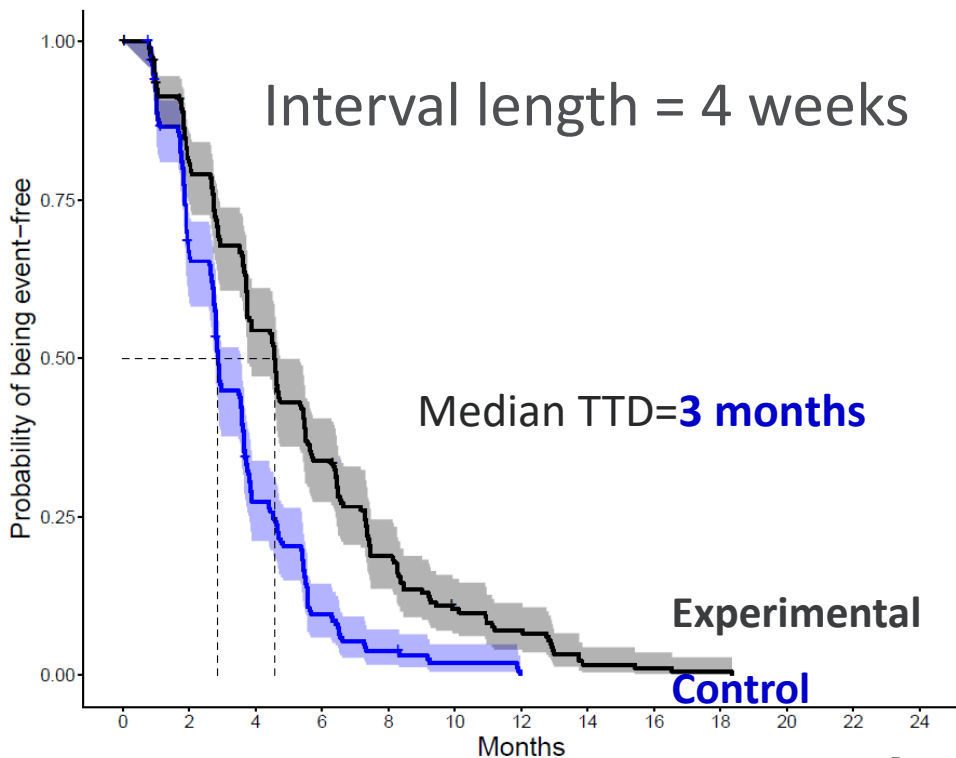
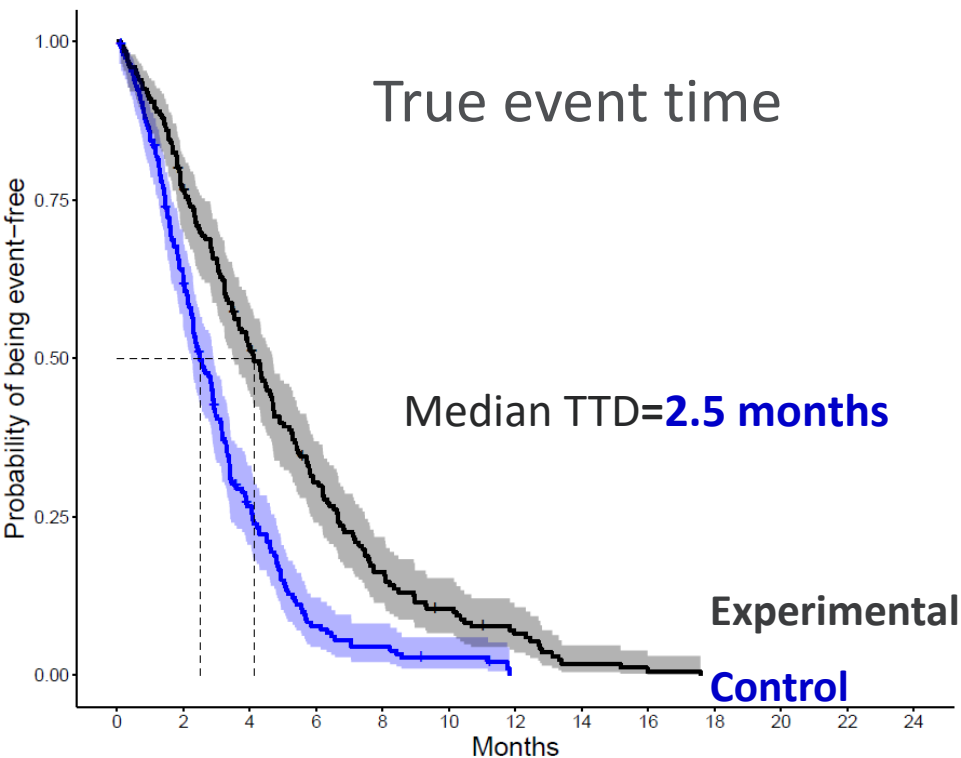
# Scenario 1 – on-treatment assessments at same frequency between arms

## Simulation methods

- > **Fixed-visit imputation** – assign event time to the visit at which it is first detected
- > Assume **no informative censoring**
  - > Only fixed data cut off:  $C_{1i} \sim U(0, 100 \text{ months})$
- > Vary the true **HR = {0.5, 0.75, 1}** favouring experimental arm
- > **HR < 1** indicates delayed deterioration in experimental arm

# Scenario 1 – on-treatment assessments at same frequency between arms

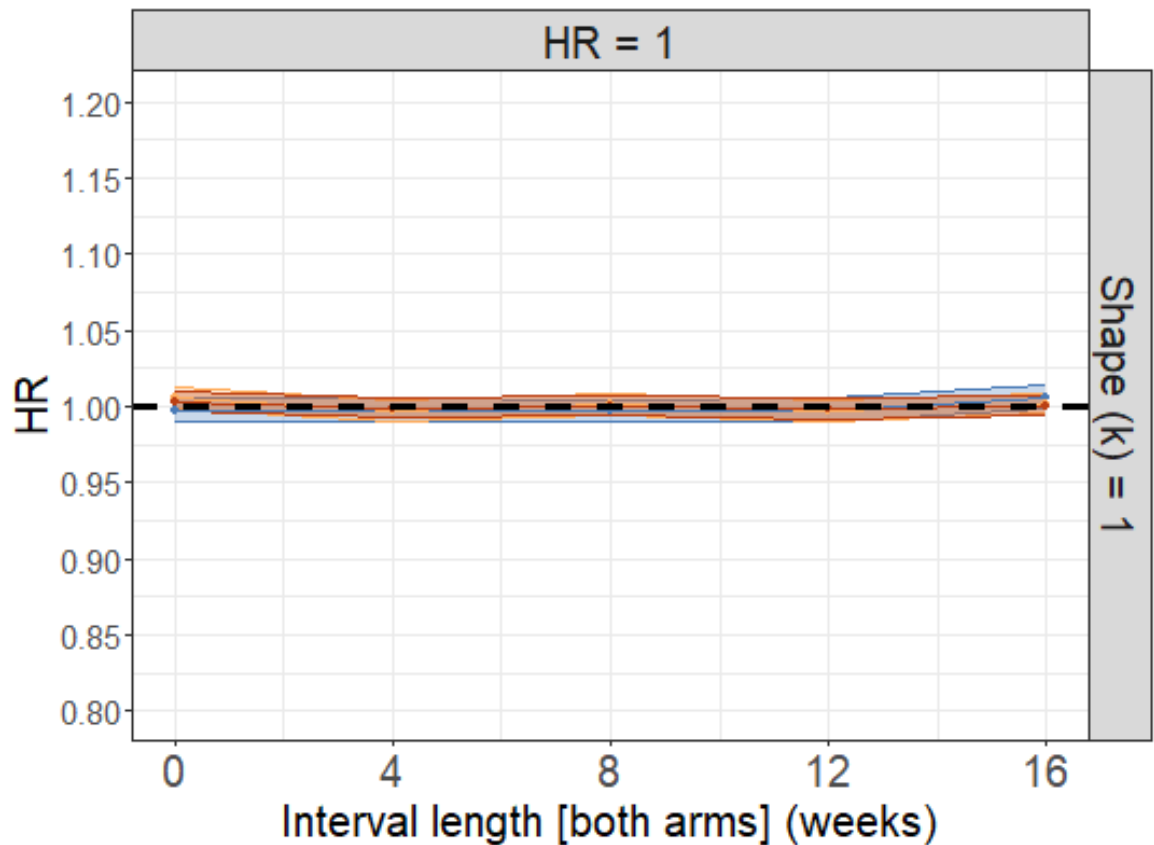
Simulation results



(HR=0.5, lambda=0.17, k=1.5)

# Scenario 1 – on-treatment assessments at same frequency between arms

Simulation results – Bias (HR)



> Log(HR),  $\beta'_{Trx}$ , averaged over multiple simulations ( $n_{sim} = 1000$ )

> Shaded areas represent Monte-Carlo 95% CI ('simulation error')

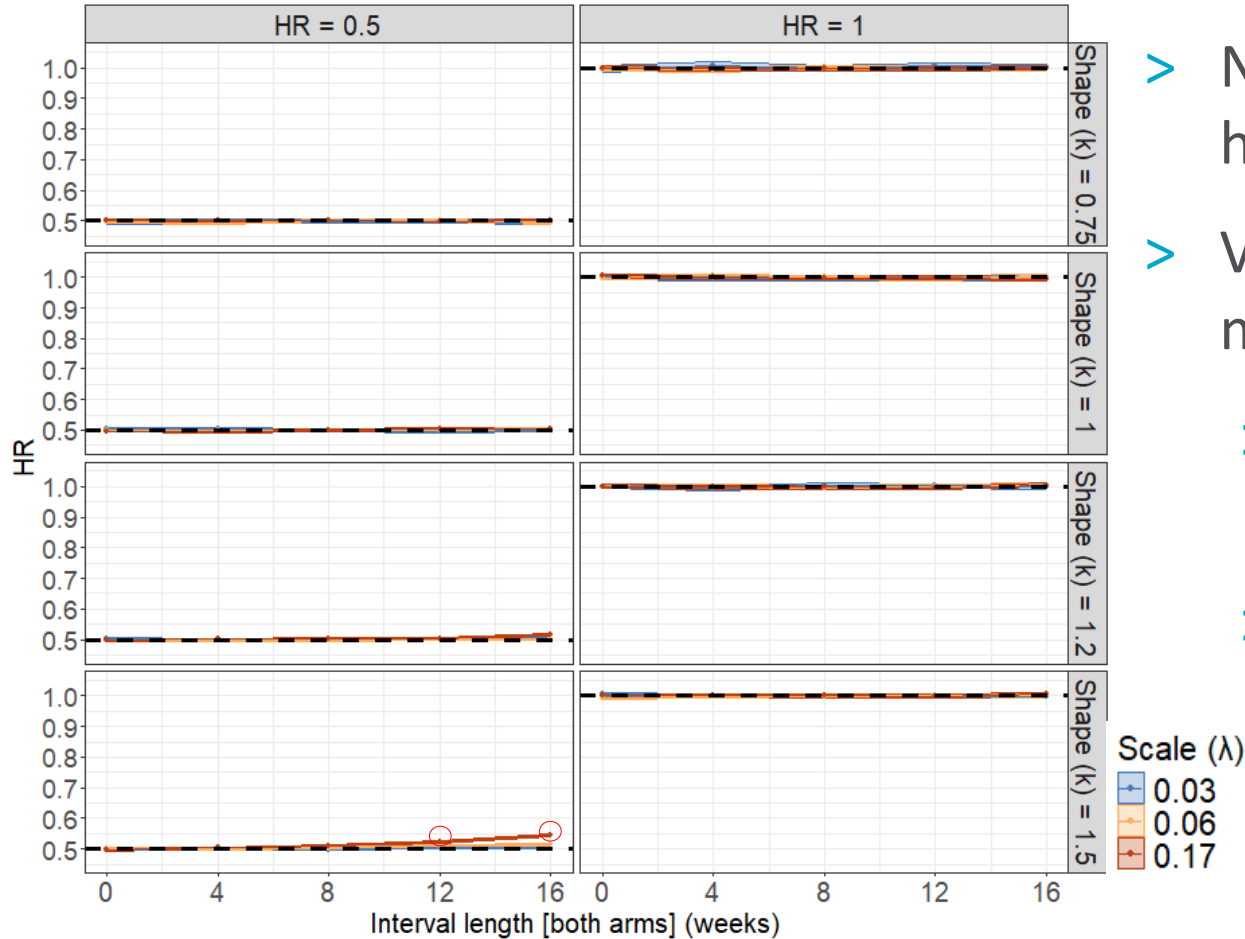
Scale ( $\lambda$ )

- 0.03
- 0.06
- 0.17

> Dotted line - - -, marks **true HR** (HR=1)

> Deviation from this line indicates **bias** of the Cox PH model

# Scenario 1 – on-treatment assessments at same frequency between arms



- > No bias across almost all hazard distributions
- > Very minimal bias in the most extreme scenario
  - > Bias is in direction of the **null**
  - > Bias increases as true treatment difference increases

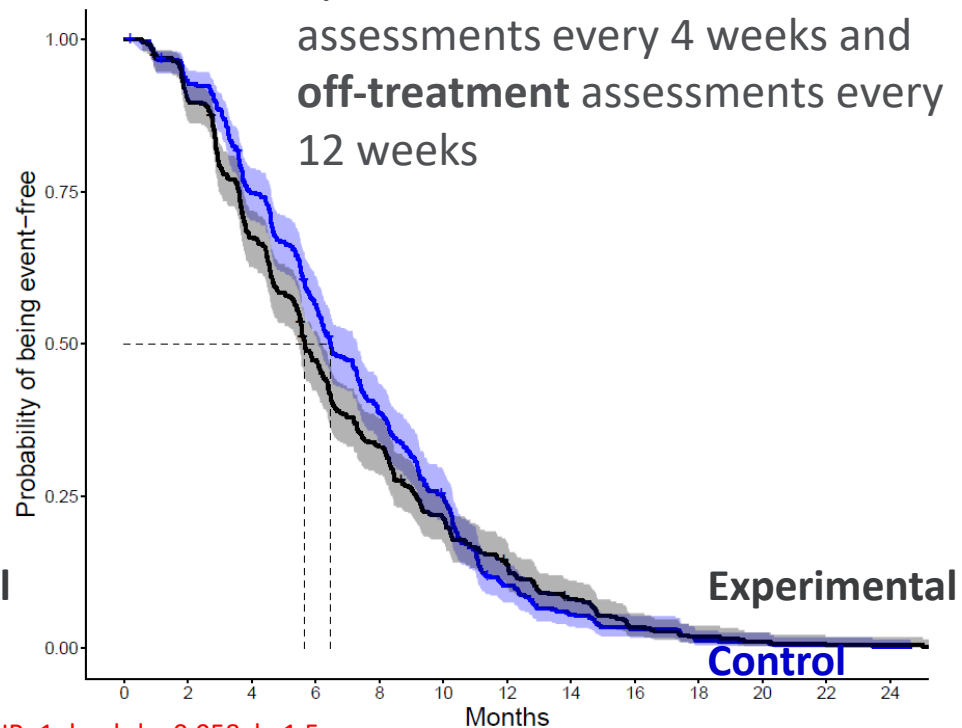
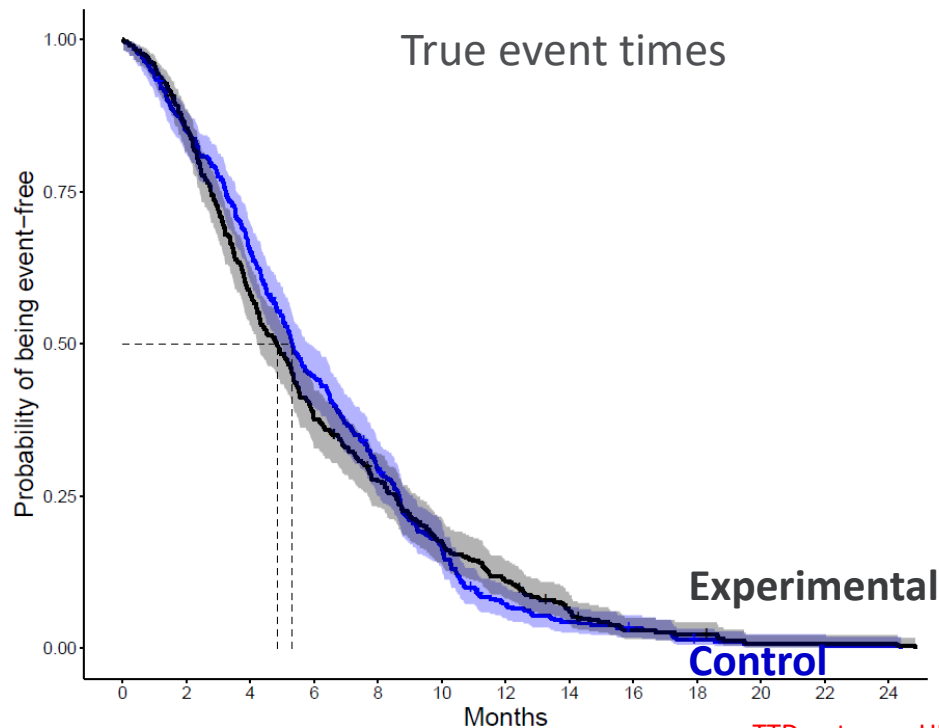
# Scenario 2 – assessment frequency reduced to every 12 weeks **post-ICE**

## Methods

- > Assume **no informative censoring**
- > Introduce **intercurrent event (ICE)** → occurrence of which triggers change in PRO assessment schedule i.e. ‘off-treatment schedule’
  - > Assume ‘off treatment’ collection fixed to be every 12 weeks **indefinitely**
- > Simulate ICE times from **exponential distribution**
- > Vary rate of ICE ( $\lambda$ ) and **ICE HR = {0.5, 0.75, 1}**
- > Repeat simulations with increasing correlation between PRO TTD and ICE times (dependence structure introduced using **Gaussian copula**)<sub>12</sub>

# Scenario 2 – assessment frequency reduced to every 12 weeks post-ICE

Simulation results

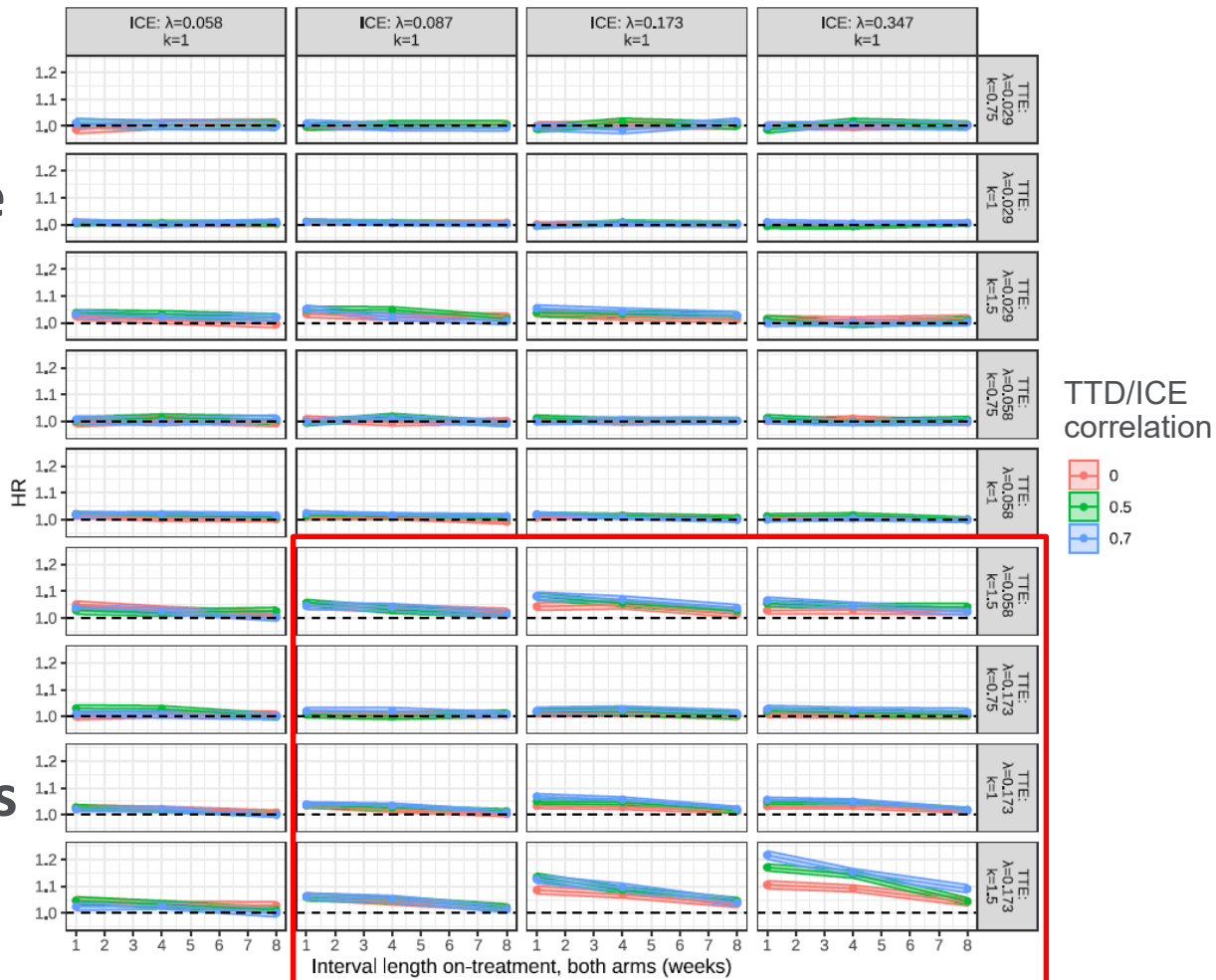


TTD outcome HR=1;  $\lambda=0.058$ ;  $k=1.5$ ;  
ICE HR=0.5;  $\lambda=0.173$

# Scenario 2 – assessment frequency reduced to every 12 weeks **post-ICE**

For simplicity:

- > Assume no difference in PRO TTD (HR=1): black dotted line - - -
- > Assume HR for ICE=0.5 (ICEs more frequent in control)
- > This results is a **non-negligible level of bias** in favour of control



# Conclusions

- > Scenario 1 shows us that Cox PH model generally robust when assessment schedule symmetric between arms, even if interval length large (16 weeks)
  - > **Kaplan-Meier estimates** will still be biased even if Cox HR is not!
- > Scenario 2 demonstrates asymmetry can be introduced through key design features related to post-ICE collection
  - > This introduces bias favouring the arm with highest rate of transition to lower frequency assessments

# Key takeaways

- > Avoid introducing any asymmetry between arms as much as possible
- > Regular PRO assessment frequency continuing **post-ICE** is the least biased approach
- > Other **informative/dependent censoring** mechanisms are likely more important to address than interval censoring
  - > Focus should remain on robust strategies for handling right-censored data not observed after ICEs

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Thank you for listening

**Any questions?**